The economic commitment of climate change

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Global projections of macroeconomic climate-change damages typic 'v co-sider impacts from average annual and national temperatures over long time in vizor s²⁻⁶. Here we use recent empirical findings from more than 1,600 regions worldwase over the past 40 years to project sub-national damages from temporatu. and recipitation, including daily variability and extremes^{7,8}. Using an empirical approa_{1,1} that provides a robust lower bound on the persistence of impacts on economic frowth, we find that the world economy is committed to an income reduction of within the next 26 years independent of future emission choic is (relation to a baseline without climate impacts, likely range of 11–29% accor nting for physical climate and empirical uncertainty). These damages already our eighthe logation costs required to limit global warming to 2 °C by sixfold over is n ar-term time frame and thereafter diverge strongly dependent on emission choices. Sommitted damages arise predominantly through changes in average ten prature, but accounting for further climatic components raises estimates by proprint and leads to stronger regional heterogeneity. Committed losses are projected for all regions except those at very high latitudes, at which reactions in temperature variability bring benefits. The largest losses are co. mitted at lower latitudes in regions with lower cumulative historical emissions and over present-day income.

Projections of the macroeconomic damage caused by ruce climate change are crucial to informing public and policy debates about a daptation, mitigation and climate justice. On the case hand, adaptation against climate impacts must be justified and ponned on the basis of an understanding of their future magnitude and antial distribution. This is also of importance in the context o mate justice¹⁰, as well as to key societal actors, including governments, contibanks and private businesses, which increasingly require the liclusion of climate risks in their macroeconomic forecas's to a dadap live decision-making 11,12. On the other hand, climate mitig tion, cy such as the Paris Climate Agreement is often evalue (ed by ba. soing the costs of its implementation against the benefits of voiding projected physical damages. This evaluation occurs 20th form 11v through cost-benefit analyses 1,4-6, as well as informally through public perception of mitigation and damage costs 3.

Project; s of fure damages meet challenges when informing these d' bate , in part cular the human biases relating to uncertainty and remediless and are raised by long-term perspectives¹⁴. Here we aim to overce such challenges by assessing the extent of economic damages from climate change to which the world is already committed by historical emissions and socio-economic inertia (the range of future emission scenarios that are considered socio-economically plausible¹⁵). Such a focus on the near term limits the large uncertainties about diverging future emission trajectories, the resulting long-term climate response and the validity of applying historically observed climate-economic relations over long timescales during which socio-technical conditions may change considerably. As such, this focus aims to simplify the communication and maximize the credibility of projected economic damages from future climate change.

In projecting the future economic damages from climate change, we make use of recent advances in climate econometrics that provide evidence for impacts on sub-national economic growth from numerous components of the distribution of daily temperature and precipitation^{3,7,8}. Using fixed-effects panel regression models to control for potential confounders, these studies exploit within-region variation in local temperature and precipitation in a panel of more than 1,600 regions worldwide, comprising climate and income data over the past 40 years, to identify the plausibly causal effects of changes in several climate variables on economic productivity^{16,17}. Specifically, macroeconomic impacts have been identified from changing daily temperature variability, total annual precipitation, the annual number of wet days and extreme daily rainfall that occur in addition to those already identified from changing average temperature^{2,3,18}. Moreover, regional heterogeneity in these effects based on the prevailing local climatic conditions has been found using interactions terms. The selection of these climate variables follows micro-level evidence for mechanisms related to the impacts of average temperatures on labour and agricultural productivity², of temperature variability on agricultural productivity and health⁷, as well as of precipitation on agricultural productivity, labour outcomes and flood damages8 (see Extended Data Table 1 for an overview, including more detailed references). References 7,8 contain a more detailed motivation for the use of these particular climate variables and provide extensive empirical tests about the robustness and nature of their effects on economic output, which are summarized in Methods. By accounting for these extra climatic variables at the sub-national level, we aim for a more comprehensive description of climate impacts with greater detail across both time and space.

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Constraining the persistence of impacts

A key determinant and source of discrepancy in estimates of the magnitude of future climate damages is the extent to which the impact of a climate variable on economic growth rates persists. The two extreme cases in which these impacts persist indefinitely or only instantaneously are commonly referred to as growth or level effects^{19,20} (see Methods section 'Empirical model specification: fixed-effects distributed lag models' for mathematical definitions). Recent work shows that future damages from climate change depend strongly on whether growth or level effects are assumed²⁰. Following refs. 2,18, we provide constraints on this persistence by using distributed lag models to test the significance of delayed effects separately for each climate variable. Notably, and in contrast to refs. 2.18, we use climate variables in their first-differenced form following ref. 3, implying a dependence of the growth rate on a change in climate variables. This choice means that a baseline specification without any lags constitutes a model prior of purely level effects, in which a permanent change in the climate has only an instantaneous effect on the growth rate^{3,19,21}. By including lags, one can then test whether any effects may persist further. This is in contrast to the specification used by refs. 2,18, in which climate variables are used without taking the first difference, implying a dependence of the growth rate on the level of climate variables. In this alternative case, the baseline specification without any lags constitutes a model prior of pure growth effects, in which a change in climate has an infinitely persistent effect on the growth rate. Consequently, including further lags in this alternative case tests whether the initial growth impact is recovered^{18,19,21}. Both of these specifications suffer from the limiting possibility that, if too few lags are included, one might falsely accept the model prior. The limitations of including a very large number of lags, including loss of data and increasing statistical uncertain with an increasing number of parameters, mean that such a possib. " is likely. By choosing a specification in which the model prior is one of level effects, our approach is therefore conservative by c'esi, avoiding assumptions of infinite persistence of climate impacts on growth and instead providing a lower bound on this persist race based on what is observable empirically (see Methods section 'E apirical relodel specification: fixed-effects distributed lag models' for . "ther 'xposition of this framework). The conservative nature such a choice is probably the reason that ref. 19 finds much greater concentrations between the impacts projected by models the first difference of climate variables, as opposed to their vels.

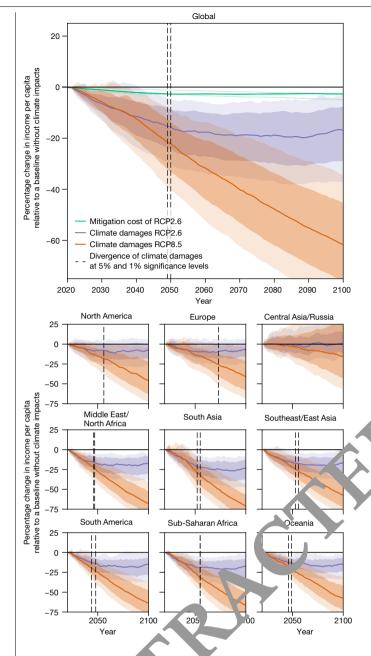
We begin our empirical analys. Sistence of climate impacts on growth using ten lage of the fire differenced climate variables in fixed-effects distriby .ea g mode.s. We detect substantial effects on economic grow that time as of up to approximately 8–10 years for the temper cure terms and up to approximately 4 years for the precipitation ter (Exter ded Data Fig. 1 and Extended Data Table 2). Furthermo evalua proy means of information criteria indicates that the inc¹ sion of all five climate variables and the use of these numbers of lags provide a preferable trade-off between best-fitting the data and including further terms that could cause overfitting, in comparison with model specifications excluding climate variables or including more or fewer lags (Extended Data Fig. 3, Supplementary Methods Section 1 and Supplementary Table 1). We therefore remove statistically insignificant terms at later lags (Supplementary Figs. 1-3 and Supplementary Tables 2-4). Further tests using Monte Carlo simulations demonstrate that the empirical models are robust to autocorrelation in the lagged climate variables (Supplementary Methods Section 2 and Supplementary Figs. 4 and 5), that information criteria provide an effective indicator for lag selection (Supplementary Methods Section 2 and Supplementary Fig. 6), that the results are robust to concerns of imperfect multicollinearity between climate variables and that including several climate variables is actually necessary to isolate their separate effects (Supplementary Methods Section 3 and Supplementary Fig. 7).

We provide a further robustness check using a restricted distributed lag model to limit oscillations in the lagged parameter estimates that may result from autocorrelation, finding that it provides similar estimates of cumulative marginal effects to the unrestricted model (Supplementary Methods Section 4 and Supplementary Figs. 8 and 9). Finally, to explicitly account for any outstanding uncertainty arising from the precise choice of the number of lags, we include empirical models with marginally different numbers of lags in the error-sampling procedure of our projection of future damages. On the basis of the lag-selection procedure (the significance of lagged terms in Extended Dan Fig. 1 and Extended Data Table 2, as well as information criteria in tended Data Fig. 3), we sample from models with eight to tan lags for a perature and four for precipitation (models shown in 5 pleme stary Figs. 1-3 and Supplementary Tables 2–4). In sum 7, 10, 14, this mr rical approach to constrain the persistence of climate; impacts on economic growth rates is conservative by design in avoiding assuraptions of infinite persistence, but nevertheless proves a sound on the extent of impact persistence that is robastion numerous tests outlined above.

Committed damages until aid-century

We combine thes implicate conomic response functions (Supplementary Figs. 1-3 and S. plementary Tables 2-4) with an ensemble of 21 climate models (see Suprementary Table 5) from the Coupled Model Intercomp Troject Phase 6 (CMIP-6)²² to project the macroeconomic dama res from these components of physical climate change (see Methods for further details). Bias-adjusted climate models that pro. 'e a highly accurate reproduction of observed climatological patter, with limited uncertainty (Supplementary Table 6) are used to oid i' troducing biases in the projections. Following a well-developed lite ature^{2,3,19}, these projections do not aim to provide a prediction of uture economic growth. Instead, they are a projection of the exogenous impact of future climate conditions on the economy relative to the baselines specified by socio-economic projections, based on the plausibly causal relationships inferred by the empirical models and assuming ceteris paribus. Other exogenous factors relevant for the prediction of economic output are purposefully assumed constant.

A Monte Carlo procedure that samples from climate model projections, empirical models with different numbers of lags and model parameter estimates (obtained by 1,000 block-bootstrap resamples of each of the regressions in Supplementary Figs. 1-3 and Supplementary Tables 2-4) is used to estimate the combined uncertainty from these sources. Given these uncertainty distributions, we find that projected global damages are statistically indistinguishable across the two most extreme emission scenarios until 2049 (at the 5% significance level; Fig. 1). As such, the climate damages occurring before this time constitute those to which the world is already committed owing to the combination of past emissions and the range of future emission scenarios that are considered socio-economically plausible¹⁵. These committed damages comprise a permanent income reduction of 19% on average globally (population-weighted average) in comparison with a baseline without climate-change impacts (with a likely range of 11-29%, following the likelihood classification adopted by the Intergovernmental Panel on Climate Change (IPCC); see caption of Fig. 1). Even though levels of income per capita generally still increase relative to those of today, this constitutes a permanent income reduction for most regions, including North America and Europe (each with median income reductions of approximately 11%) and with South Asia and Africa being the most strongly affected (each with median income reductions of approximately 22%; Fig. 1). Under a middle-of-the road scenario of future income development (SSP2, in which SSP stands for Shared Socio-economic Pathway), this corresponds to global annual damages in 2049 of 38 trillion in 2005 international dollars (likely range of 19-59 trillion 2005 international dollars). Compared with empirical specifications that assume pure growth or pure level effects,



our preferred specification that provides a robust lower bound on the extent of climate in pact pers. ance produces damages between these two extreme assumptions (Extended Data Fig. 3).

Damares a ready outweigh mitigation costs

the lamages to which the world is committed over the next 25 year estimates of the mitigation costs required to achieve the Paris Climate Agreement. Taking estimates of mitigation costs from the three integrated assessment models (IAMs) in the IPCC AR6 database²³ that provide results under comparable scenarios (SSP2 baseline and SSP2-RCP2.6, in which RCP stands for Representative Concentration Pathway), we find that the median committed climate damages are larger than the median mitigation costs in 2050 (six trillion in 2005 international dollars) by a factor of approximately six (note that estimates of mitigation costs are only provided every 10 years by the IAMs and so a comparison in 2049 is not possible). This comparison simply aims to compare the magnitude of future damages against mitigation costs, rather than to conduct a formal cost-benefit analysis of transitioning from one emission path to another. Formal

Fig. 1 | The commitment and divergence of economic climate damages versus mitigation costs. Estimates of the projected reduction in income per capita from changes in all climate variables based on empirical models of climate impacts on economic output with a robust lower bound on their persistence (Extended Data Fig. 1) under a low-emission scenario compatible with the 2 °C warming target and a high-emission scenario (SSP2-RCP2.6 and SSP5-RCP8.5, respectively) are shown in purple and orange, respectively. Shading represents the 34% and 10% confidence intervals reflecting the likely and very likely ranges, respectively (following the likelihood classification adopted by the IPCC), having estimated uncertainty from a Monte Carlo procedure, which samples the uncertainty from the choice of anysical climate models, empirical models with different numbers of lags at a poor scarped estimates of the regression parameters shown in Supplementa Figs. 1-3. Vertical dashed lines show the time at which the climate damages the two emission scenarios diverge at the 5% and 1% significa alevels ased on the distribution of differences between emission pharic rising from the uncertainty sampling discussed above. Note that uncertainty of the difference of the two scenarios is smaller than the courbined uncertainty of the two respective scenarios because sample of the certainty (climate model and empirical model choice, as well as node arameter bootstrap) are consistent across the two emission scenar os, hence to divergence of damages occurs while the uncertainty bour is o. etwo separate damage scenarios still overlap. Estimates of global mitigat a costs from the three IAMs that provide results for the SSP2 basen and SSP2 RCP2.6 scenario are shown in light green in the top panel, with the radian of these estimates shown in bold.

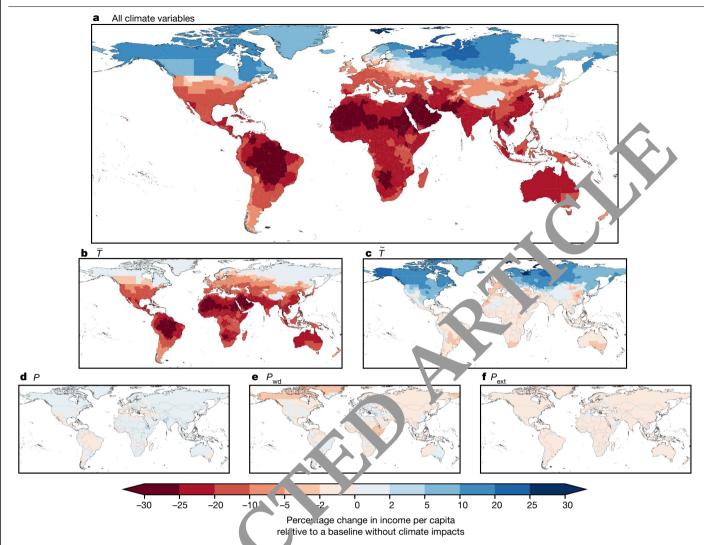
cost-beneficenally estypically find that the net benefits of mitigation only emerge: fter 2050 (ref. 5), which may lead some to conclude that pny, al damages from climate change are simply not large enough to out reigh mitigation costs until the second half of the century. Our aple comparison of their magnitudes makes clear that damages are ctually already considerably larger than mitigation costs and the 'elayed emergence of net mitigation benefits results primarily from the fact that damages across different emission paths are indistinguishable until mid-century (Fig. 1).

Although these near-term damages constitute those to which the world is already committed, we note that damage estimates diverge strongly across emission scenarios after 2049, conveying the clear benefits of mitigation from a purely economic point of view that have been emphasized in previous studies^{4,24}. As well as the uncertainties assessed in Fig. 1, these conclusions are robust to structural choices. such as the timescale with which changes in the moderating variables of the empirical models are estimated (Supplementary Figs. 10 and 11), as well as the order in which one accounts for the intertemporal and international components of currency comparison (Supplementary Fig. 12; see Methods for further details).

Damages from variability and extremes

Committed damages primarily arise through changes in average temperature (Fig. 2). This reflects the fact that projected changes in average temperature are larger than those in other climate variables when expressed as a function of their historical interannual variability (Extended Data Fig. 4). Because the historical variability is that on which the empirical models are estimated, larger projected changes in comparison with this variability probably lead to larger future impacts in a purely statistical sense. From a mechanistic perspective, one may plausibly interpret this result as implying that future changes in average temperature are the most unprecedented from the perspective of the historical fluctuations to which the economy is accustomed and therefore will cause the most damage. This insight may prove useful in terms of guiding adaptation measures to the sources of greatest damage.

Nevertheless, future damages based on empirical models that consider changes in annual average temperature only and exclude the other climate variables constitute income reductions of only 13% in 2049



arising separately from changes in annual mean temperature (\mathbf{b}), daily temperature variability (\mathbf{c}), total annual precipitation (\mathbf{d}), the annual number of wet days (>1 mm) (\mathbf{e}) and extreme daily rainfall (\mathbf{f}) (see Methods for further definitions). Data on national administrative boundaries are obtained from the GADM database version 3.6 and are freely available for academic use (https://gadm.org/).

(Extended Data Fig. 3), likely and 5–21%). This suggests that accounting for the other components of the distribution of temperature and precipitation rate and mages by nearly 50%. This increase arises through the further lar ages that these climatic components cause, but also because their inclusion reveals a stronger negative economic response average temperatures (Extended Data Fig. 5b). The latter finding is consistent with our Monte Carlo simulations, which suggest that the magnitude of the effect of average temperature on economic growth is underestimated unless accounting for the impacts of other correlated climate variables (Supplementary Fig. 7).

In terms of the relative contributions of the different climatic components to overall damages, we find that accounting for daily temperature variability causes the largest increase in overall damages relative to empirical frameworks that only consider changes in annual average temperature (4.9 percentage points, likely range 2.4–8.7 percentage points, equivalent to approximately 10 trillion international dollars). Accounting for precipitation causes smaller increases in overall damages, which are—nevertheless—equivalent to approximately 1.2 trillion international dollars: 0.01 percentage points (-0.37–0.33 percentage points), 0.34 percentage points (0.07–0.90 percentage points) and

0.36 percentage points (0.13–0.65 percentage points) from total annual precipitation, the number of wet days and extreme daily precipitation, respectively. Moreover, climate models seem to underestimate future changes in temperature variability 25 and extreme precipitation 26,27 in response to anthropogenic forcing as compared with that observed historically, suggesting that the true impacts from these variables may be larger.

The distribution of committed damages

The spatial distribution of committed damages (Fig. 2a) reflects a complex interplay between the patterns of future change in several climatic components and those of historical economic vulnerability to changes in those variables. Damages resulting from increasing annual mean temperature (Fig. 2b) are negative almost everywhere globally, and larger at lower latitudes in regions in which temperatures are already higher and economic vulnerability to temperature increases is greatest (see the response heterogeneity to mean temperature embodied in Extended Data Fig. 1a). This occurs despite the amplified warming projected at higher latitudes²⁸, suggesting that regional heterogeneity

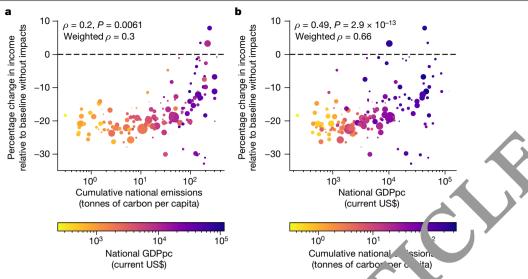


Fig. 3 | The injustice of committed climate damages by cumulative historical emissions and income. Estimates of the median projected change in national income per capita across emission scenarios (RCP2.6 and RCP8.5) as well as climate model, empirical model and model parameter uncertainty in the year in which climate damages diverge at the 5% level (2049, as identified in Fig. 1) are plotted against cumulative national emissions per capita in 2020 (from the

Global Carbon Project) and co. red by nat onal income per capita in 2020 (from the World Bank) in **a** and vice ersa in **b**. In each panel, the size of each scatter point is weight with enational population in 2020 (from the World Bank). Inset number a indicate the Spearman's rank correlation ρ and P-values for a hypothesis test versus appothesis is of no correlation, as well as the Spearman's rank correlative eighted by national population.

in economic vulnerability to temperature changes outweighs heterogeneity in the magnitude of future warming (Supplementary Fig. 13a). Economic damages owing to daily temperature variability (Fig. 2c) exhibit a strong latitudinal polarisation, primarily reflecting he physical response of daily variability to greenhouse forcing in which increases in variability across lower latitudes (and Europe) co. ".st decreases at high latitudes²⁵ (Supplementary Fig. 13b). These t temperature terms are the dominant determinants of the pattern of overall damages (Fig. 2a), which exhibits a strong rolarity with damages across most of the globe except at the highest northern lar audes. Future changes in total annual precipitation mainly bring economic benefits except in regions of drying, such as the 'edit tranean and central South America (Fig. 2d and Supplementary Fig. 13c), but these benefits are opposed by changes in the jurge of wet days, which produce damages with a similar pattern of opposite sign (Fig. 2e and Supplementary Fig. 13d). By cor trast changes in extreme daily rainfall produce damages in all regions, field, the intensification of daily rainfall extremes over g' bai land eas^{29,30} (Fig. 2f and Supplementary Fig. 13e).

The spatial distribution of primitted damages implies considerable injustice a'ong wo dimensions: culpability for the historical emissions that he cause I climate change and pre-existing levels of socio-eco mic we rr. Spearman's rank correlations indicate that commi^r ed d mages are significantly larger in countries with smaller historical mula ve emissions, as well as in regions with lower current income per wita (Fig. 3). This implies that those countries that will suffer the most from the damages already committed are those that are least responsible for climate change and which also have the least resources to adapt to it.

To further quantify this heterogeneity, we assess the difference in committed damages between the upper and lower quartiles of regions when ranked by present income levels and historical cumulative emissions (using a population weighting to both define the quartiles and estimate the group averages). On average, the quartile of countries with lower income are committed to an income loss that is 8.9 percentage points (or 61%) greater than the upper quartile (Extended Data Fig. 6), with a likely range of 3.8–14.7 percentage points across the uncertainty sampling of our damage projections (following the likelihood classification adopted by the IPCC). Similarly, the quartile of cou, ies with lower historical cumulative emissions are committed to an income loss that is 6.9 percentage points (or 40%) greater than the oper quartile, with a likely range of 0.27–12 percentage points. These particle in climate impacts 31-33 the context of the damages to which the world is already committed by historical emissions and socio-economic inertia.

Contextualizing the magnitude of damages

The magnitude of projected economic damages exceeds previous literature estimates^{2,3}, arising from several developments made on previous approaches. Our estimates are larger than those of ref. 2 (see first row of Extended Data Table 3), primarily because of the facts that sub-national estimates typically show a steeper temperature response (see also refs. 3,34) and that accounting for other climatic components raises damage estimates (Extended Data Fig. 5). However, we note that our empirical approach using first-differenced climate variables is conservative compared with that of ref. 2 in regard to the persistence of climate impacts on growth (see introduction and Methods section 'Empirical model specification: fixed-effects distributed lag models'), an important determinant of the magnitude of long-term damages 19,21. Using a similar empirical specification to ref. 2, which assumes infinite persistence while maintaining the rest of our approach (sub-national data and further climate variables), produces considerably larger damages (purple curve of Extended Data Fig. 3). Compared with studies that do take the first difference of climate variables^{3,35}, our estimates are also larger (see second and third rows of Extended Data Table 3). The inclusion of further climate variables (Extended Data Fig. 5) and a sufficient number of lags to more adequately capture the extent of impact persistence (Extended Data Figs. 1 and 2) are the main sources of this difference, as is the use of specifications that capture nonlinearities in the temperature response when compared with ref. 35. In summary, our estimates develop on previous studies by incorporating the latest data and empirical insights^{7,8}, as well as in providing a robust empirical lower bound on the persistence of impacts on economic growth, which constitutes a middle ground between the extremes of the growth-versus-levels debate^{19,21} (Extended Data Fig. 3).

Compared with the fraction of variance explained by the empirical models historically (<5%), the projection of reductions in income of 19% may seem large. This arises owing to the fact that projected changes in climatic conditions are much larger than those that were experienced historically, particularly for changes in average temperature (Extended Data Fig. 4). As such, any assessment of future climate-change impacts necessarily requires an extrapolation outside the range of the historical data on which the empirical impact models were evaluated. Nevertheless, these models constitute the most state-of-the-art methods for inference of plausibly causal climate impacts based on observed data. Moreover, we take explicit steps to limit out-of-sample extrapolation by capping the moderating variables of the interaction terms at the 95th percentile of the historical distribution (see Methods). This avoids extrapolating the marginal effects outside what was observed historically. Given the nonlinear response of economic output to annual mean temperature (Extended Data Fig. 1 and Extended Data Table 2), this is a conservative choice that limits the magnitude of damages that we project. Furthermore, back-of-the-envelope calculations indicate that the projected damages are consistent with the magnitude and patterns of historical economic development (see Supplementary Discussion Section 5).

Missing impacts and spatial spillovers

Despite assessing several climatic components from which economic impacts have recently been identified 3,7,8, this assessment of aggregate climate damages should not be considered comprehensive. Important channels such as impacts from heatwaves³¹, sea-level rise³⁶, tropical cyclones³⁷ and tipping points^{38,39}, as well as non-market damages such as those to ecosystems⁴⁰ and human health⁴¹, are not considered in these estimates. Sea-level rise is unlikely to be feasibly incorparated into empirical assessments such as this because historical sea-leve variability is mostly small. Non-market damages are inherent! intracta within our estimates of impacts on aggregate monetary atput and estimates of these impacts could arguably be considered as extra to those identified here. Recent empirical work sug ,ests that accounting for these channels would probably raise estimates of these committed damages, with larger damages continuing a prise in the global south^{31,36-42}.

Moreover, our main empirical analysis loes. explicitly evaluate the potential for impacts in local regions to produce effects that 'spill over' into other regions. Such refect may further mitigate or amplify the impacts we estimate, for example, in a spanies relocate production from one affected region o anothe. rifimpacts propagate along supply chains. The current ate. *ure indicates that trade plays a substantial role in propagating pillover capets 43,44, making their assessment at the sub-national level challenging without available data on sub-national trade dependen ... Stuc'es accounting for only spatially adjacent neighbour dicate at negative impacts in one region induce further negativ impacts in neighbouring regions 45-48, suggesting that our projected da. ges a. probably conservative by excluding these effects. In Supplement Fig. 14, we assess spillovers from neighbouring regions using a spatial-lag model. For simplicity, this analysis excludes temporal lags, focusing only on contemporaneous effects. The results show that accounting for spatial spillovers can amplify the overall magnitude, and also the heterogeneity, of impacts. Consistent with previous literature, this indicates that the overall magnitude (Fig. 1) and heterogeneity (Fig. 3) of damages that we project in our main specification may be conservative without explicitly accounting for spillovers. We note that further analysis that addresses both spatially and trade-connected spillovers, while also accounting for delayed impacts using temporal lags, would be necessary to adequately address this question fully. These approaches offer fruitful avenues for further research but are beyond the scope of this manuscript, which primarily aims to explore the impacts of different climate conditions and their persistence.

Policy implications

We find that the economic damages resulting from climate change until 2049 are those to which the world economy is already committed and that these greatly outweigh the costs required to mitigate emissions in line with the 2 °C target of the Paris Climate Agreement (Fig. 1). This assessment is complementary to formal analyses of the net costs and benefits associated with moving from one emission path to another, which typically find that net benefits of magation only emerge in the second half of the century⁵. Our simple conparison of the magnitude of damages and mitigation costs make. Year that this is primarily because damages are indistinguishable acros missions scenarios—that is, committed—until mid-cent. v(Fig. 1 and that they are actually already much larger than microtion starfor simplicity. and owing to the availability of data, we compare dar .ages to mitigation costs at the global level. Regional est nates of mitigation costs may shed further light on the nation. 'ncen. or mitigation to which our results already hint, of receance for international climate policy. Although these damages 2. ommittee from a mitigation perspective, adaptation may provide an operation or continuity to reduce them. Moreover, the strong diverger of damas is after mid-century reemphasizes the clear benefit of maigation from a purely economic perspective, as highlighted in prous scudies 1,4,6,24.

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Any methods, additional references, Nature Portfolio reporting summanes, urce data, extended data, supplementary information, acknowledgen. nts, peer review information; details of author contributions d conpeting interests; and statements of data and code availability are vailable at https://doi.org/10.1038/s41586-024-07219-0.

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Methods

Historical climate data

Historical daily 2-m temperature and precipitation totals (in mm) are obtained for the period 1979–2019 from the W5E5 database. The W5E5 dataset comes from ERA-5, a state-of-the-art reanalysis of historical observations, but has been bias-adjusted by applying version 2.0 of the WATCH Forcing Data to ERA-5 reanalysis data and precipitation data from version 2.3 of the Global Precipitation Climatology Project to better reflect ground-based measurements⁴⁹⁻⁵¹. We obtain these data on a $0.5^{\circ} \times 0.5^{\circ}$ grid from the Inter-Sectoral Impact Model Intercomparison Project (ISIMIP) database. Notably, these historical data have been used to bias-adjust future climate projections from CMIP-6 (see the following section), ensuring consistency between the distribution of historical daily weather on which our empirical models were estimated and the climate projections used to estimate future damages. These data are publicly available from the ISIMIP database. See refs. 7,8 for robustness tests of the empirical models to the choice of climate data reanalysis products.

Future climate data

Daily 2-m temperature and precipitation totals (in mm) are taken from 21 climate models participating in CMIP-6 under a high (RCP8.5) and a low (RCP2.6) greenhouse gas emission scenario from 2015 to 2100. The data have been bias-adjusted and statistically downscaled to a common half-degree grid to reflect the historical distribution of daily temperature and precipitation of the WSE5 dataset using the trend-preserving method developed by the ISIMIP 50,52 . As such, the climate model data reproduce observed climatological patterns exceptionally well $(^{\rm C}$ upplementary Table 5). Gridded data are publicly available from the SIMIP database.

Historical economic data

Historical economic data come from the DOSF 'atabase f subnational economic output⁵³. We use a recent revision to the DOSE dataset that provides data across 83 countries, 1,660 rub-national regions with varying temporal coverage from 1960 to ub-national units constitute the first administrati for example, states for the USA and provinces. China. Data come from measures of gross region product per capita (GRPpc) or income per capita in local curr noies reflecting the values reported in national statistical agencie vea 20ksd, in some cases, academic literature. We follow provious literature 3,7,8,54 and assess real subnational output per copit. v first converting values from local currencies to US dollars to accomplor diverging national inflationary tendencies and then account for US inflation using a US deflator. Alternatively, on gight first account for national inflation and then convert been cultiplies. Supplementary Fig. 12 demonstrates that our collusions are consistent when accounting for price changes in the reversed order, although the magnitude of estimated damages varies. The documentation of the DOSE dataset for further discussion of these choices. Conversions between currencies are conducted using exchange rates from the FRED database of the Federal Reserve Bank of St. Louis⁵⁵ and the national deflators from the World Bank56.

Future socio-economic data

Baseline gridded gross domestic product (GDP) and population data for the period 2015–2100 are taken from the middle-of-the-road scenario SSP2 (ref. 15). Population data have been downscaled to a half-degree grid by the ISIMIP following the methodologies of refs. 57,58, which we then aggregate to the sub-national level of our economic data using the spatial aggregation procedure described below. Because current methodologies for downscaling the GDP of the SSPs use downscaled population to do so, per-capita estimates of GDP with a

realistic distribution at the sub-national level are not readily available for the SSPs. We therefore use national-level GDP per capita (GDPpc) projections for all sub-national regions of a given country, assuming homogeneity within countries in terms of baseline GDPpc. Here we use projections that have been updated to account for the impact of the COVID-19 pandemic on the trajectory of future income, while remaining consistent with the long-term development of the SSPs. The choice of baseline SSP alters the magnitude of projected climate damages in monetary terms, but when assessed in terms of parentage change from the baseline, the choice of socio-economic sce. Trio is inconsequential. Gridded SSP population data are inational-likel GDPpc data are publicly available from the ISIMIP of a base. Sub-national estimates as used in this study are a anable in the code and data replication files.

Climate variables

Following recent literature², we calculate an array of climate variables for which substant an apacts of macroeconomic output have been identified empirically, poorted by further evidence at the micro level for plaus le under ying mechanisms. See refs. 7,8 for an extensive mediation for the use of these particular climate variables and for detaile empirical tests on the nature and robustness of their effection econo. Ic output. To summarize, these studies have found evid nec independent impacts on economic growth rates from annual verage temperature, daily temperature variability, total annual precipitation, the annual number of wet days and extreme daily rain. Assessments of daily temperature variability were motivated by evicence of impacts on agricultural output and human health, well as macroeconomic literature on the impacts of volatility on gre vth when manifest in different dimensions, such as government pending, exchange rates and even output itself⁷. Assessments of precipitation impacts were motivated by evidence of impacts on agricultural productivity, metropolitan labour outcomes and conflict, as well as damages caused by flash flooding8. See Extended Data Table 1 for detailed references to empirical studies of these physical mechanisms. Marked impacts of daily temperature variability, total annual precipitation, the number of wet days and extreme daily rainfall on macroeconomic output were identified robustly across different climate datasets, spatial aggregation schemes, specifications of regional time trends and error-clustering approaches. They were also found to be robust to the consideration of temperature extremes $^{7.8}$. Furthermore, these climate variables were identified as having independent effects on economic output^{7,8}, which we further explain here using Monte Carlo simulations to demonstrate the robustness of the results to concerns of imperfect multicollinearity between climate variables (Supplementary Methods Section 2), as well as by using information criteria (Supplementary Table 1) to demonstrate that including several lagged climate variables provides a preferable trade-off between optimally describing the data and limiting the possibility of overfitting.

We calculate these variables from the distribution of daily, d, temperature, $T_{x,d}$, and precipitation, $P_{x,d}$, at the grid-cell, x, level for both the historical and future climate data. As well as annual mean temperature, $\overline{T}_{x,y}$, and annual total precipitation, $P_{x,y}$, we calculate annual, y, measures of daily temperature variability, $\widetilde{T}_{x,y}$:

$$\widetilde{T}_{x,y} = \frac{1}{12} \sum_{m=1}^{12} \sqrt{\frac{1}{D_m} \sum_{d=1}^{D_m} (T_{x,d,m,y} - \overline{T}_{x,m})^2},$$
 (1)

the number of wet days, $Pwd_{x,y}$:

$$Pwd_{x,y} = \sum_{d=1}^{D_y} H(P_{x,d} - 1 \text{ mm})$$
 (2)

and extreme daily rainfall:

$$Pext_{x,y} = \sum_{d=1}^{D_y} H(P_{x,d} - P99.9_x) \times P_{x,d},$$
 (3)

in which $T_{x,d,m,y}$ is the grid-cell-specific daily temperature in month m and year y, $\overline{T}_{x,m,y}$ is the year and grid-cell-specific monthly, m, mean temperature, D_m and D_y the number of days in a given month m or year y, respectively, H the Heaviside step function, 1 mm the threshold used to define wet days and $P99.9_x$ is the 99.9th percentile of historical (1979-2019) daily precipitation at the grid-cell level. Units of the climate measures are degrees Celsius for annual mean temperature and daily temperature variability, millimetres for total annual precipitation and extreme daily precipitation, and simply the number of days for the annual number of wet days.

We also calculated weighted standard deviations of monthly rainfall totals as also used in ref. 8 but do not include them in our projections as we find that, when accounting for delayed effects, their effect becomes statistically indistinct and is better captured by changes in total annual rainfall.

Spatial aggregation

We aggregate grid-cell-level historical and future climate measures, as well as grid-cell-level future GDPpc and population, to the level of the first administrative unit below national level of the GADM database, using an area-weighting algorithm that estimates the portion of each grid cell falling within an administrative boundary. We use this as our baseline specification following previous findings that the effect of area or population weighting at the sub-national level is negligible ⁷⁸.

Empirical model specification: fixed-effects distributed la models

Following a wide range of climate econometric literatur (1660), we upanel regression models with a selection of fixed effect, and time trends to isolate plausibly exogenous variation with which praximize confidence in a causal interpretation of the effects of climate on economic growth rates. The use of region fixed effects, in, accounts for unobserved time-invariant differences between regions, such as prevailing climatic norms and growth rate, and go historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects, ing to historical and geopolitical factors. The use of yearly fixed effects and geopolitical factors are used to historical and geopolitical factors. The use of yearly fixed effects are used to historical and geopolitical factors. The use of yearly fixed effects are used to historical and geopolitical factors. The use of yearly fixed effects are used to historical and geopolitical factors. The use of yearly fixed effects are used to historical and geopolitical factors. The use of yearly fixed effects are used to historical and geopolitical factors.

The persistence of climate pacts on economic growth rates is a key determinant of the long-term magnitude of damages. Methods for inferring the external figure at stance in impacts on growth rates have typically used in the variables to evaluate the presence of delayed effects that dynamics $^{2.18}$. For example, consider starting from a model in thick a climate condition, $C_{r,y}$ (for example, annual mean temperature infects the growth rate, $\Delta \lg rp_{r,y}$ (the first difference of the logarithm of gross regional product) of region r in year y:

$$\Delta \operatorname{Igrp}_{r,y} = \mu_r + \eta_y + k_r y + \alpha C_{r,y} + \varepsilon_{r,y}, \tag{4}$$

which we refer to as a 'pure growth effects' model in the main text. Typically, further lags are included,

$$\Delta \operatorname{Igrp}_{r,y} = \mu_r + \eta_y + k_r y + \sum_{L=0}^{NL} \alpha_L C_{r,y-L} + \varepsilon_{r,y}, \tag{5}$$

and the cumulative effect of all lagged terms is evaluated to assess the extent to which climate impacts on growth rates persist. Following ref. 18, in the case that,

$$\sum_{L=0}^{NL} \alpha_L < 0 \text{ for } \alpha_0 < 0 \text{ or } \sum_{L=0}^{NL} \alpha_L > 0 \text{ for } \alpha_0 > 0,$$
 (6)

the implication is that impacts on the growth rate persist up to NL years after the initial shock (possibly to a weaker or a stronger extent), whereas if

$$\sum_{L=0}^{NL} \alpha_L = 0, \tag{7}$$

then the initial impact on the growth rate is r covered at . NL years and the effect is only one on the level of output. 'owever we note that such approaches are limited by the fact that, when reluding an insufficient number of lags to detect a recovery of the growth rates, one may find equation (6) to be satisfied and incorrectly assume that a change in climatic conditions at the partial with rate indefinitely. In practice, given a limited record of r torical data, including too few lags to confidently concluding an infinitely persistent impact on the growth rate is likely, particula. Every the long timescales over which future climate damage are often projected r To avoid this issue, we instead begin out that r are often projected r To avoid this issue, we instead begin out the level of a climate variable, r Cry:

$${}^{1}\alpha r p_{y} = \mu_r + \eta_y + k_r y + \alpha C_{r,y} + \varepsilon_{r,y}. \tag{8}$$

Given the non-stationarity of the level of output, we follow the literature and estimate such an equation in first-differenced form as,

$$\Delta \operatorname{Igrp}_{r,y} = \mu_r + \eta_y + k_r y + \alpha \Delta C_{r,y} + \varepsilon_{r,y}, \tag{8}$$

hich we refer to as a model of 'pure level effects' in the main text. This model constitutes a baseline specification in which a permanent change in the climate variable produces an instantaneous impact on the growth rate and a permanent effect only on the level of output. By including lagged variables in this specification,

$$\Delta \operatorname{Igrp}_{r,y} = \mu_r + \eta_y + k_r y + \sum_{l=0}^{NL} \alpha_l \Delta C_{r,y-l} + \varepsilon_{r,y}, \tag{9}$$

we are able to test whether the impacts on the growth rate persist any further than instantaneously by evaluating whether $\alpha_L > 0$ are statistically significantly different from zero. Even though this framework is also limited by the possibility of including too few lags, the choice of a baseline model specification in which impacts on the growth rate do not persist means that, in the case of including too few lags, the framework reverts to the baseline specification of level effects. As such, this framework is conservative with respect to the persistence of impacts and the magnitude of future damages. It naturally avoids assumptions of infinite persistence and we are able to interpret any persistence that we identify with equation (9) as a lower bound on the extent of climate impact persistence on growth rates. See the main text for further discussion of this specification choice, in particular about its conservative nature compared with previous literature estimates, such as refs. 2,18.

We allow the response to climatic changes to vary across regions, using interactions of the climate variables with historical average (1979–2019) climatic conditions reflecting heterogenous effects identified in previous work. Following this previous work, the moderating variables of these interaction terms constitute the historical average of either the variable itself or of the seasonal temperature difference, \hat{T}_r , or annual mean temperature, \overline{T}_r , in the case of daily temperature variability and extreme daily rainfall, respectively.

The resulting regression equation with *N* and *M* lagged variables, respectively, reads:

$$\Delta \operatorname{Igrp}_{r,y} = \mu_{r} + \eta_{y} + k_{r}y + \sum_{L=0}^{N} (\alpha_{1,L} \Delta \overline{T}_{r,y-L} + \alpha_{2,L} \Delta \overline{T}_{r,y-L} \times \overline{T}_{r})$$

$$+ \sum_{L=0}^{N} (\alpha_{3,L} \Delta \widetilde{T}_{r,y-L} + \alpha_{4,L} \Delta \widetilde{T}_{r,y-L} \times \hat{T}_{r})$$

$$+ \sum_{L=0}^{M} (\alpha_{5,L} \Delta P_{r,y-L} + \alpha_{6,L} \Delta P_{r,y-L} \times P_{r})$$

$$+ \sum_{L=0}^{M} (\alpha_{7,L} \Delta \operatorname{Pwd}_{r,y-L} + \alpha_{8,L} \Delta \operatorname{Pwd}_{r,y-L} \times \operatorname{Pwd}_{r})$$

$$+ \sum_{L=0}^{M} (\alpha_{9,L} \Delta \operatorname{Pext}_{r,y-L} + \alpha_{10,L} \Delta \operatorname{Pext}_{r,y-L} \times \overline{T}_{r}) + \epsilon_{r,y}$$

$$(10)$$

in which Δ lgrp_{r,y} is the annual, regional GRPpc growth rate, measured as the first difference of the logarithm of real GRPpc, following previous work^{2,3,7,8,18,19}. Fixed-effects regressions were run using the fixest package in R (ref. 61).

Estimates of the coefficients of interest $\alpha_{i,L}$ are shown in Extended Data Fig. 1 for N=M=10 lags and for our preferred choice of the number of lags in Supplementary Figs. 1–3. In Extended Data Fig. 1, errors are shown clustered at the regional level, but for the construction of damage projections, we block-bootstrap the regressions by region 1,000 times to provide a range of parameter estimates with which to sample the projection uncertainty (following refs. 2,31).

Spatial-lag model

In Supplementary Fig. 14, we present the results from a spatial-lag model that explores the potential for climate impacts to 'spill c, c' into spatially neighbouring regions. We measure the distance be week centroids of each pair of sub-national regions and construct $s_{\rm p}$ —all lags that take the average of the first-differenced climate variables at their interaction terms over neighbouring regions that are a distances of 0–500, 500–1,000, 1,000–1,500 and 1,500–200 km (spa. Mags, 'SL', 1 to 4). For simplicity, we then assess a spatial-lag model without temporal lags to assess spatial spillovers of continuous climate impacts. This model takes the form:

$$\Delta \operatorname{Igrp}_{r,y} = \mu_{r} + \eta_{y} + k_{r}y + \sum_{SL=0}^{N} (\alpha_{1,SL} \Delta \overline{T}_{r} + \alpha_{2,SL}, \overline{T}_{r-SL,y} \times \overline{T}_{r-SL})$$

$$+ \sum_{SL=0}^{N} (\alpha_{3,SL} \Delta \widetilde{T}_{r-SL,y} + \sum_{SL,y} k_{r-SL,y} \times r_{r-SL})$$

$$+ \sum_{SL=0}^{N} (\alpha_{5,SL} \wedge P_{r-SL,y} + \sum_{SL,y} k_{r-SL,y} \times P_{r-SL})$$

$$+ \sum_{SL=0}^{N} (\alpha_{7,SL} \wedge Pwd_{r-SL,y} + \alpha_{8,SL} \Delta Pwd_{r-SL,y} \times Pwd_{r-SL})$$

$$+ \sum_{SL=0}^{N} (\alpha_{7,SL} \wedge Pwd_{r-SL,y} + \alpha_{8,SL} \Delta Pwd_{r-SL,y} \times Pwd_{r-SL})$$

$$+ \sum_{SL=0}^{N} (\alpha_{7,SL} \wedge Pwd_{r-SL,y} + \alpha_{10,SL} \Delta Pext_{r-SL,y} \times \overline{T}_{r-SL}) + \epsilon_{r,y},$$
(11)

in which SL indicates the spatial lag of each climate variable and interaction term. In Supplementary Fig. 14, we plot the cumulative marginal effect of each climate variable at different baseline climate conditions by summing the coefficients for each climate variable and interaction term, for example, for average temperature impacts as:

ME =
$$\sum_{SL=0}^{N} (\alpha_{1,SL} + \alpha_{2,SL} \overline{T}_{r-SL}).$$
 (12)

These cumulative marginal effects can be regarded as the overall spatially dependent impact to an individual region given a one-unit shock to a climate variable in that region and all neighbouring regions at a given value of the moderating variable of the interaction term.

Constructing projections of economic damage from future climate change

We construct projections of future climate damages by applying the coefficients estimated in equation (10) and shown in Supplementary Tables 2-4 (when including only lags with statistically significant effects in specifications that limit overfitting; see Supplementary Methods Section 1) to projections of future climate change from the MIP-6 models. Year-on-year changes in each primary climate variation of interest are calculated to reflect the year-to-year variations used in the empirical models. 30-year moving averages of the moderating variables of the interaction terms are calculated to reflect the ong-te maverage of climatic conditions that were used for the mode at ing variables in the empirical models. By using moving averages in the projections, we account for the changing vulnerability oclimate shocks based on the evolving long-term conditions (2 pple). ...ary Figs. 10 and 11 show that the results are robust to the precise choice of the window of this moving average). Although seclima evariables are not differenced, the fact that the bias-aquiste. I limate models reproduce observed climatological pattern across regions for these moderating variables very accurately (pp) mentary Table 6) with limited spread across models (<3%) precises the possibility that any considerable bias or uncertain sintroduc aby this methodological choice. However, we impose caps of moderating variables at the 95th percentile at which they vare observed in the historical data to prevent extrapolation of the marginal effects outside the range in which the regressions were timated. This is a conservative choice that limits the magnitude of our amage projections.

Times series of primary climate variables and moderating climate variables are then combined with estimates of the empirical model arameters to evaluate the regression coefficients in equation (10), producing a time series of annual GRPpc growth-rate reductions for a given emission scenario, climate model and set of empirical model parameters. The resulting time series of growth-rate impacts reflects those occurring owing to future climate change. By contrast, a future scenario with no climate change would be one in which climate variables do not change (other than with random year-to-year fluctuations) and hence the time-averaged evaluation of equation (10) would be zero. Our approach therefore implicitly compares the future climate-change scenario to this no-climate-change baseline scenario.

The time series of growth-rate impacts owing to future climate change in region r and year y, $\delta_{r,y}$, are then added to the future baseline growth rates, $\pi_{r,y}$ (in log-diff form), obtained from the SSP2 scenario to yield trajectories of damaged GRPpc growth rates, $\rho_{r,y}$. These trajectories are aggregated over time to estimate the future trajectory of GRPpc with future climate impacts:

GRPpc_{r,Y} = GRPpc_{r,2020}
$$\sum_{y=2020}^{Y} \rho_{r,y}$$

= GRPpc_{r,2020} $\sum_{y=2020}^{Y} (1 + \pi_{r,y} + \delta_{r,y})$, (13)

in which $GRPpc_{r,y=2020}$ is the initial log level of GRPpc. We begin damage estimates in 2020 to reflect the damages occurring since the end of the period for which we estimate the empirical models (1979–2019) and to match the timing of mitigation-cost estimates from most IAMs (see below).

For each emission scenario, this procedure is repeated 1,000 times while randomly sampling from the selection of climate models, the selection of empirical models with different numbers of lags (shown in Supplementary Figs. 1-3 and Supplementary Tables 2-4) and bootstrapped estimates of the regression parameters. The result is an ensemble of future GRPpc trajectories that reflect uncertainty from

both physical climate change and the structural and sampling uncertainty of the empirical models.

Estimates of mitigation costs

We obtain IPCC estimates of the aggregate costs of emission mitigation from the AR6 Scenario Explorer and Database hosted by IIASA²³. Specifically, we search the AR6 Scenarios Database World v1.1 for IAMs that provided estimates of global GDP and population under both a SSP2 baseline and a SSP2-RCP2.6 scenario to maintain consistency with the socio-economic and emission scenarios of the climate damage projections. We find five IAMs that provide data for these scenarios, namely, MESSAGE-GLOBIOM 1.0, REMIND-MAgPIE 1.5, AIM/GCE 2.0, GCAM 4.2 and WITCH-GLOBIOM 3.1. Of these five IAMs, we use the results only from the first three that passed the IPCC vetting procedure for reproducing historical emission and climate trajectories. We then estimate global mitigation costs as the percentage difference in global per capita GDP between the SSP2 baseline and the SSP2-RCP2.6 emission scenario. In the case of one of these IAMs, estimates of mitigation costs begin in 2020, whereas in the case of two others, mitigation costs begin in 2010. The mitigation cost estimates before 2020 in these two IAMs are mostly negligible, and our choice to begin comparison with damage estimates in 2020 is conservative with respect to the relative weight of climate damages compared with mitigation costs for these two IAMs.

Data availability

Data on economic production and ERA-5 climated at a are blicly available at https://doi.org/10.5281/zenodo.4 681306 (ref. 62) and https://www.ecmwf.int/en/forecasts/datasets/recallysis-d/tasets/era5, respectively. Data on mitigation costs are publicly ole at https:// data.ene.iiasa.ac.at/ar6/#/downloads. ssed climate and economic data, as well as all other necessary cata to, eproduction of the results, are available at the publication v https://doi.org/10.5281/ zenodo.10562951 (ref. 63).

Code availability

All code necessar, for reproduction of the results is available at the public reposite vhttv://doi.org/10.5281/zenodo.10562951 (ref. 63).

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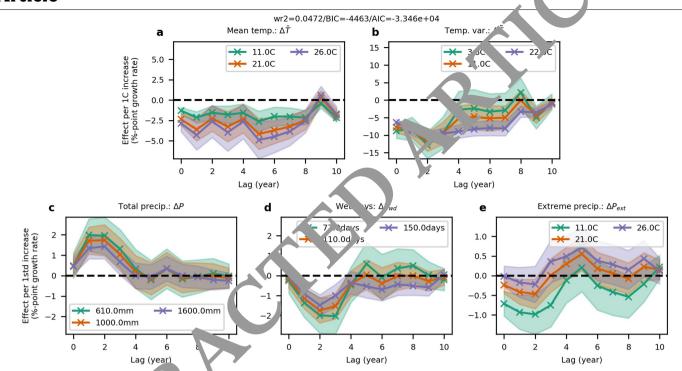
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Additional information

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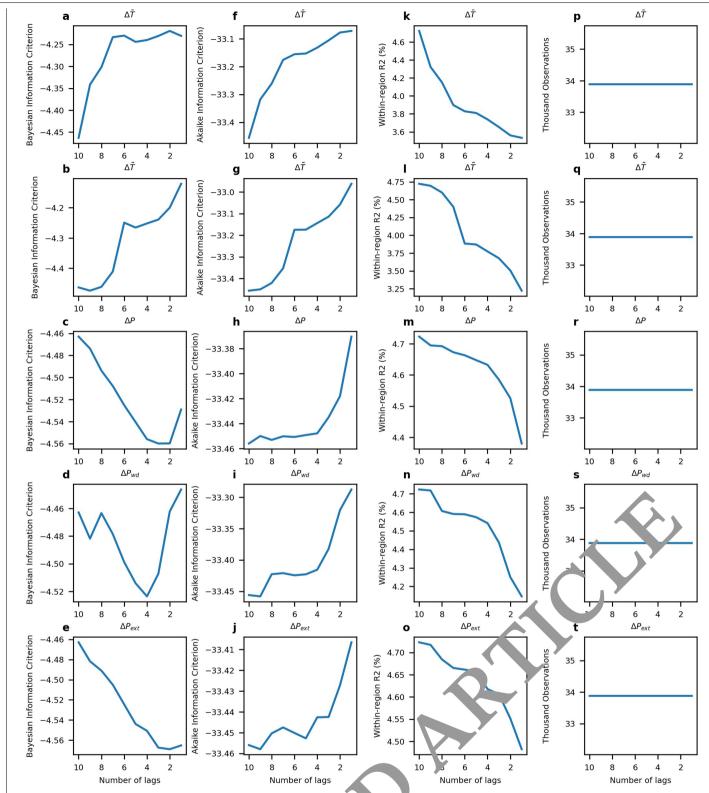
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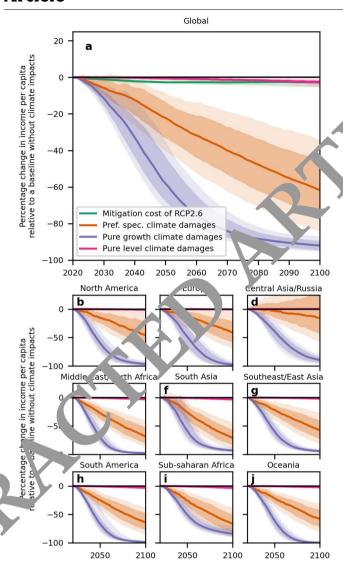
Extended Data Fig. 1 | Constraining the Persistence of historical climate impacts on economic growth rates The reads of a panel-based fixed-effects distributed lag model for the effects on a pula mean temperature (a), daily temperature variability (**), to a lannual precipitation (c), the number of wet days (d) and extreme (hill) precipitation (e) on sub-national economic growth rates. Point estimates show the effects of a 1 °C or one standard deviation increase (for temperature an 1 precipitation variables, respectively) at the lower quartile media. Individual precipitation variables, respectively at the initial singuistic (procedule) are not cumulative effects). Climate variables are used in the rest-differenced form (see main text for discussion) and the

moderating climate variables are the annual mean temperature, seasonal temperature difference, total annual precipitation, number of wet days and annual mean temperature, respectively, in panels $\mathbf{a}-\mathbf{e}$ (see Methods for further discussion). Error bars show the 95% confidence intervals having clustered standard errors by region. The within-region \mathbf{R}^2 , Bayesian and Akaike information criteria for the model are shown at the top of the figure. This figure shows results with ten lags for each variable to demonstrate the observed levels of persistence, but our preferred specifications remove later lags based on the statistical significance of terms shown above and the information criteria shown in Extended Data Fig. 2. The resulting models without later lags are shown in Supplementary Figs. 1–3.

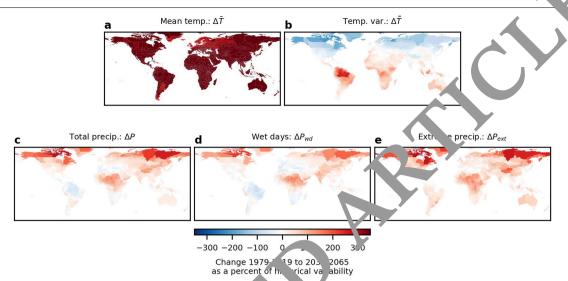


Extended Data Fig. 2 | Incremental lag-selection procedure using information criteria and within-region R^2 . Starting from a panel-base of itself-effects distributed lag model estimating the effects of climate on economic growth using the real historical data (as in equation (4)) with tenlings for an climate variables (as shown in Extended Data Fig. 1), lags are in a centrally removed for one climate variable at a time. The resulting Baye diana. Akaike information criteria are shown in $\mathbf{a} - \mathbf{e}$ and $\mathbf{f} - \mathbf{j}$, respectively, and the with degion R^2 and number of observations in $\mathbf{k} - \mathbf{o}$ and $\mathbf{p} - \mathbf{t}$, respectively. Different rows

ow the results when removing lags from different climate variables, ordered free cop to bottom as annual mean temperature, daily temperature variability, or all annual precipitation, the number of wet days and extreme annual precipitation. Information criteria show minima at approximately four lags for precipitation variables and ten to eight for temperature variables, indicating that including these numbers of lags does not lead to overfitting. See Supplementary Table 1 for an assessment using information criteria to determine whether including further climate variables causes overfitting.

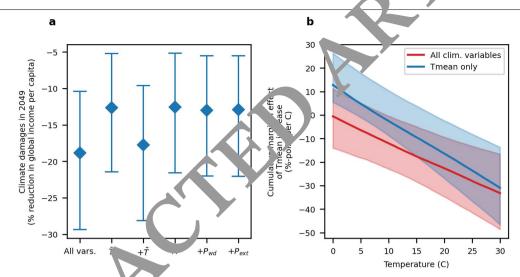


Extended Data Fig. 3 | Damages in our preferred specification that provides a robust lower bound on the persistence of climate impacts on economic growth versus damages in specifications of pure growth or pure level effects. Estimates of future damages as shown in Fig. 1 but under the emission scenario RCP8.5 for three separate empirical specifications: in orange our preferred specification, which provides an empirical lower bound on the persistence of climate impacts on economic growth rates while avoiding assumptions of infinite persistence (see main text for further discussion); in purple a specification of 'pure growth effects' in which the first difference of climate variables is not taken and no lagged climate variables are included (the baseline specification of ref. 2); and in pink a specification of 'pure level effects' in which the first difference of climate variables is taken but no lagged terms are included.



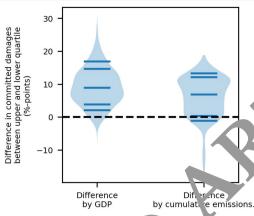
Extended Data Fig. 4 | **Climate changes in different variables as a function of historical interannual variability.** Changes in each climate variable or part of from 1979–2019 to 2035–2065 under the high-emission scenario SCros RCP expressed as a percentage of the historical variability of each medical. Historical variability is estimated as the standard deviation of each detended to limate variable over the period 1979–2019 during which the engine all models are

ide ... ied (detrending is appropriate because of the inclusion of region-specific inear time trends in the empirical models). See Supplementary Fig. 13 for changes expressed in standard units. Data on national administrative boundaries are obtained from the GADM database version 3.6 and are freely available for academic use (https://gadm.org/).



Extended Data Fig. 5 | Contribution lifferen climate variables to overall committed damages. a, Climate dimage in 2049 then using empirical models that account for all climate variables. Thing, annual mean temperature only or changes in both annual mian temperature and one other climate variable (daily temperature variable) that could be annual precipitation, the number of wet days and extreme dails on cipitation, respectively). **b**, The cumulative marginal

effects of an increase in annual mean temperature of 1 $^{\circ}$ C, at different baseline temperatures, estimated from empirical models including all climate variables or annual mean temperature only. Estimates and uncertainty bars represent the median and 95% confidence intervals obtained from 1,000 block-bootstrap resamples from each of three different empirical models using eight, nine or ten lags of temperature terms.



Extended Data Fig. 6 | The difference i...com. "ted damages between the upper and lower quartiles of count." when raked by GDP and cumulative historical emissions. Quartiles a defin. "usin a population weighting, as are the average committed damage across in quartile group. The violin plots indicate the distribution difference between quartiles across the two extreme emission scenarios (RCP and RCP8.5) and the uncertainty sampling procedure outlined in thoos, while accounts for uncertainty arising from the choice of lags in the entirical models, uncertainty in the empirical model parameter estimates, as well the climate model projections. Bars indicate the median, is well as the 10th and 90th percentiles and upper and lower sixths of the distribution reflecting the very likely and likely ranges following the likelihood class.

Extended Data Table 1 | A summary of several physical mechanisms that p a suc, underlie the impact of the different climate variables on macroeconomic growth, with references to empirical avidence

Climate variable	Physical mechanisms	References
Average annual temperature	Labour productivity and supply; agriculture productivity	Dasgupta et al. (2021) ⁶² ;Lobell et al. (2013) ⁶³ , Zhao et al. (2017) ⁶⁴
Daily temperature variability	Agricultural productivity; physical realth; menu balth	Wheeler et al. $(2000)^{65}$, Rowhani et al. $(2011)^{66}$, Ceglar et al. $(2016)^{67}$; Shi et al. $(2015)^{68}$; Xue et al. $(2019)^{69}$
Total annual precipitation	Agricultural productivity; metropolitan labour utcomes; conflict	Liang et al. (2017) $^{70};$ Desbreaux et al. (2019) $^{71};$ Damania et al. (2020) 72
Number of wet days	Tri el disruption	Lacking
Extreme daily precipitation	Flood dameusruption	Davenport et al. (2021) ⁷³ , Dave et al. (2021) ⁷⁴

This summary is not intended to be an exhause list of at mechanisms or references. In the case of most climate variables, several plausible physical mechanisms supported by empirical evidence exist. The only exception here; the number of we days, for which plausible mechanisms are listed but empirical evidence does not yet exist (as far as the authors are aware). The use of the number of wet days in the main employable therefore guided primarily by the empirical evidence indicating robust impacts on economic growth. References 64–76 in the table.

Extended Data Table 2 | Regression results for the historical effects of different climate valiable on sub-national economic growth rates in the period 1979–2019

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Variable	Formula	Lag 0	Lag 1	Lag 2	Lag 3	Lag 4	Lag 5		Lag	Lag 8	Lag 9	Lag 10
Annual mean	$\Delta \bar{T}_{r,y}$	-0.17	-0.57	-0.78	-0.23	-0.79	-0.96	-0.2	-0.71	-1.8**	-1.1*	-2.5***
temperature		(0.32)	(0.5)	(0.54)	(0.57)	(0.57)	(0.65)	7.67	(0.73)	(0.63)	(0.48)	(0.35)
	$\Delta \bar{T}_{r,y}.\bar{T}_r$	-0.0011***	-0.0014**	-0.00072	-0.0015**	-0.00072	-0.0015**	-0. 7***	-0.0012**	-0.00029	0.00065*	0.00029
		(0.00029)	(0.00049)	(0.00051)	(0.00055)	(0.00047)	(0.0005)	(0.000	(0.00047)	(0.0004)	(0.00032)	(0.00023)
Daily temp.	$\Delta \tilde{T}_{r,y}$	-9.3***	-8.1***	-13***	-9.3***	-1.5	-1.?	2.4	-1.8	3.3	-5.3**	-0.34
variability	.,,	(1.3)	(2)	(2.3)	(2.7)	(3.1)	(3.2	(3)	(2.9)	(2.5)	(1.9)	(1.3)
,	$\Delta \tilde{T}_{r,y}.\hat{T}_r$	0.0013*	-0.0003	0.0013	-0.00011	-0.0034**	-0.0032 *	-0.0025*	-0.0029*	-0.003**	0.00079	-0.00037
	—,y	(0.00054)	(0.00083)	(0.00087)	(0.0011)	(0 001	(0.0013)	(0.0012)	(0.0012)	(0.0011)	(0.00078)	(0.00057)
		(0.0000.)	(0.0000)	(0.0000.)	(0.00)	, , , ,	(0.0010)	(0.00.12)	(0.00.12)	(0.0011)	(0.00070)	(0.0000.)
Total annual	$\Delta P_{r,y}$	0.002	0.0094***	0.009***	0.0068**	9021	-0.0012	0.0013	-0.001	-0.0001	0.0012	0.0005
precipitation	r,y	(0.0016)	(0.002)	(0.0023)	(0.002	(0. 24)	(0.0025)	(0.0025)	(0.0023)	(0.0021)	(0.0019)	(0.0015)
p. co.p.ianion	$\Delta P_{r,y}.P_r$	-1.4e-09	-2.6e-08**	-2.1e-08*	-2.6° 08**	-1.4e	6.3e-09	4.6e-10	6.4e-09	-1.2e-09	-1.3e-08	-9.4e-09
	r,y r	(6.9e-09)	(8.5e-09)	(9.7e-09)	(9. 09)		(1e-08)	(1e-08)	(9.3e-09)	(8.6e-09)	(7.8e-09)	(6.1e-09)
		(0.00 00)	(0.00 00)	(0000)		(,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,	(,	(10 00)	(0.00 00)	(0.00 00)	(1.00 00)	(0110 00)
Annual no.	$\Delta Pwd_{r,y}$	-0.028	-0.12**	-0.17**	-0.2**	-0.038	0.12	0.037	0.079	0.1*	0.045	-0.03
wet days		(0.038)	(0.043)	(0.05)	(0.055)	(0.052)	(0.065)	(0.068)	(0.058)	(0.048)	(0.04)	(0.032)
wor dayo	$\Delta Pwd_{r,y}.Pwd_r$	1.5e-06	4.1e-06	4.5⊾ 06	2e-06*	9.6e-07	-1e-05*	-5.5e-06	-7.1e-06	-9e-06**	-5.6e-06*	2.2e-06
		(2.3e-06)	(2.7e-06)	2-06)	(5 736)	(3.4e-06)	(4.1e-06)	(4.1e-06)	(3.6e-06)	(3e-06)	(2.5e-06)	(1.9e-06)
		(2.00 00)	(2.70 00)	00)	(00)	(0.10 00)	(0 00)	((0.00 00)	(00 00)	(2.00 00)	(1.00 00)
Precipitation	$\Delta Pext_{r,y}$	-0.023***	-0.028***	-0.029***	-0.029**	-0.01	-0.0032	-0.013	-0.017	-0.019*	-0.013	0.0054
extremes	,y	(0.0053)	(0.0073)	(0.0084)	(0.0094)	(0.0098)	(0.01)	(0.011)	(0.01)	(0.0093)	(0.0079)	(0.0052)
onti otti o	$\Delta Pext_{r,y}.\bar{T}_r$	8.8e-06***	9,6e-06**	0.000	1.4e-05**	7.7e-06	6.5e-06	8e-06	8.8e-06	8.7e-06	8.1e-06*	-1.2e-06
		(2.5e-06)	26)	(4e-06)	(4.6e-06)	(4.7e-06)	(4.9e-06)	(4.9e-06)	(4.8e-06)	(4.5e-06)	(4e-06)	(2.7e-06)
R^2	0.291	(2.00 00)	,	(.5 55)	(55)	(5 55)	(55)	(00)	(00)	((.0 00)	(30 00)
wR^2	0.0472											
BIC	-4.46e+03		T.	•								
AIC	-3.35e+04											
N	34855											
	0.000											

Numbers show the point estimate, or the effect a each climate variable and their interaction term on sub-national economic growth rates (in percentage points), having estimated equation (4) with ten lags for each climate, variable, that is, each table entry denotes a specific regression coefficient α_{XL} of the same model as indicated in equation (4)). Standard errors are shown in parentheses and *, ** and denote signature at the 5%, 1% and 0.1% levels, respectively, having clustered standard errors by region. Formulas for climate variables and their interaction terms are denoted as a equation (4). Note that an interpretation of the significance of the effects of a given climate variable requires an assessment of both the coefficient of the climate variable itself as well as a interaction term. Extended Data Fig. 1 provides the opportunity for such an interpretation by plotting the estimated marginal effects with confidence intervals. The R², within-region P^2 (the R^2 also the emporal dimension), Akaike information criterion (AIC), Bayesian information criterion (BIC) and number of observations are also shown.

Extended Data Table 3 | A comparison of the magnitude of estimated economic damage from future climate change across recent panel-based empirical studies

Study	Resolution	Number of climate variables	Baseline specification of growth- in level-effects	Nr inber of lags	Damages by 2100 under RCP8.5
Burke et al. $(2015)^2$	National	One	Growth	None	25%
Kahn et al. $(2019)^{35}$	National	One	' evel	Four	7.2%
Kalkuhl & Wenz (2020) ³	Sub-national	One	Level	One	14.2%
This study	Sub-national	Five	Level	Eight-ten/four	61.6%

All studies use fixed-effects panel regressions. The first four columns describe differences in the underlying data and empirical specification. The third column shows the nature of the baseline specification without lags with regards to grow or level effects (see main text for further discussion). The last column compares projections of future economic damage under RCP8.5 by 2100 as reported by the respective study.